



Search for the Smallest Random Forest



able to see that $(\partial/\partial\beta)\pi(\beta; k, c) = c$

$$\begin{aligned} \log(P\{M_i|y_i\}) &= -\frac{\partial}{\partial\beta}\log(P\{y_i\}) \\ &+ \sum_j \frac{\partial}{\partial\beta}\log[\pi(\beta; y_{ij})] \end{aligned}$$

the null hypothesis that $\beta = 0$, we have

$$\begin{aligned} \frac{\partial}{\partial\beta}\log[\pi(\beta; y_{ij}, 0)P\{dd|M_{ij}\}] \\ = [1 - \gamma(0; y_{ij}, 1) - \gamma(0; \end{aligned}$$

$$\frac{\partial}{\partial\beta}\log P\{y_i\}|_{\beta=0} = \sum_j [1 -$$

venience, we drop the two irrelevant

$$\begin{aligned} \log(P\{M_i|y_i\})|_{\beta=0} &= \sum_j [1 - \gamma(y_{ij}) - \gamma \\ &= \sum_j \frac{1 - \gamma(y_{ij}) - \gamma}{P\{M_{ij}\}} \end{aligned}$$

the coefficient of linkage disequilibrium

$$P\{AA\} - P\{dd, AA\} - P\{AA\}P\{DE$$

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Outline

- Background
- Goal
- Key idea
- Method
- Simulation
- Application

Background

Random forests have emerged as one of the most commonly used nonparametric statistical methods in many scientific areas, particularly in analysis of high throughput genomic data.

Background

- A general practice in using random forests is to generate a sufficiently large number of trees, although it is subjective as to how large is sufficient.
- Furthermore, random forests are viewed as a “black-box” because of its sheer size.



Goal

- Explore whether it is possible to find a common ground between a forest and a single tree
 - retain the easy interpretability of the tree-based methods
 - avoid the problems that the tree-based methods suffer from.
- Does a forest have to be large, or how small can a forest be?

Key idea

- Shrink the forest with two objectives
 - maintain a similar (or even better) level of prediction accuracy
 - reduce the number of the trees in the forest to a manageable level

Method

- Three measures are considered to determine the importance of a tree in a forest
 - by prediction
 - by similarity
 - by restricted similarity

Method

- “by prediction” method

- focuses on the prediction

- A tree can be removed if its removal from the forest has the minimal impact on the overall prediction accuracy.

$$P\{M_i|y_i\} = \frac{1}{P\{y_i\}} \prod_j [P\{y_{ij}|c_{ij} = 0\}] P\{c_{ij} = 0\}$$
$$= \frac{P\{M_i\}}{P\{y_i\}} \prod_j [\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\}]$$

$$P\{c_{ij} = 0\} = P\{y_{ij} = k | c_{ij} = 0\} = \gamma(\beta; k, c)$$
$$K - 1, \gamma(\beta, 0, c) = 0, \text{ and } \gamma(\beta, K, c) = 1. \text{ Note that}$$

$$P\{y_i\} = \prod_j [P\{y_{ij}|c_{ij} = 0\}] P\{c_{ij} = 0\}$$
$$= \prod_j [\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\}]$$

to see that $(\partial/\partial\beta)\pi(\beta; k, c) = c - \gamma(\beta; k, c)$

$$\log(P\{M_i|y_i\}) = -\frac{\partial}{\partial\beta} \log(P\{y_i\})$$
$$+ \sum_j \frac{\partial}{\partial\beta} \log[\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\}]$$

the null hypothesis that $\beta = 0$, we have

$$\frac{\partial}{\partial\beta} \log[\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\}]$$
$$= [1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)]$$

$$\frac{\partial}{\partial\beta} \log P\{y_i\}|_{\beta=0} = \sum_j [1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)]$$

convenience, we drop the two irrelevant terms

$$\log(P\{M_i|y_i\})|_{\beta=0} = \sum_j [1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)]$$
$$= \sum_j \frac{1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)}{P\{M_{ij}\}}$$

the coefficient of linkage disequilibrium

$$D(AA) = P\{dd, AA\} - P\{AA\}P\{DD\}$$

Method

- “by prediction” method

- For tree T in forest F , calculate the prediction accuracy of forest $F_{(-T)}$ that excludes T .
- $\Delta_{(-T)}$ represents the difference in prediction accuracy between F and $F_{(-T)}$.
- The tree with the smallest $\Delta_{(-T)}$ is the least important one and hence subject to removal.

Method

- “by similarity” method

- is based on the similarity between two trees.
- A tree can be removed if it is “similar” to other trees in the forest.

Method

- “by similarity” method

- The correlation of the predicted outcomes by two trees gives rise to a similarity between the two trees.
- For tree T , the average of its similarities with all trees in $F_{(-T)}$, denoted by ρ_T , reflects the overall similarity between T and $F_{(-T)}$.
- The tree with the highest ρ_T is the most similar to the trees in $F_{(-T)}$ and hence subject to removal.

Method

- “by restricted similarity” method
 - is based on the weighted similarity between two trees.
 - A tree can be removed if it is “similar” to other trees in the forest.

Method

- “by restricted similarity” method

- Evaluate the pairwise similarity of two trees in forest F , according to their predicted outcomes.
- Select the pair of trees being most similar.
- Calculate ρ_T for the two trees and the one with higher ρ_T is subject to removal.
- Distribute the weight of T to all other trees in $F_{(-T)}$, proportional to the pairwise similarity in ρ_T .

Method

- Select the optimal size sub-forest

- Let $h(i), i=1, \dots, N_f-1$, denote the performance trajectory of a sub-forest of i trees

- N_f is the size of the original random forest.

- If we have only one realization of $h(i)$, we select the optimal size sub-forest by maximizing $h(i)$ over $i=1, \dots, N_f-1$.

- If we have multiple realizations of $h(i)$, we select the optimal size sub-forest by using the 1-se rule.

- The size of this smallest sub-forest is called the critical point of the performance trajectory.

Simulation Designs

- Simulation Designs

- For each data set, we generated 500 observations, each of which has one response variable and 30 predictors from Bernoulli distribution with success probability of 0.5.
- Chose \mathcal{V} of the 30 variables to determine the response variable.

$$y = \begin{cases} 1, & \text{if } \sum_{i=1}^v X_i / v + \sigma > 0.5, \\ 0 & \text{Otherwise.} \end{cases}$$

- Where σ is a random variable following the normal distribution with mean zero and variance .
- Considered two choices for \mathcal{V} (5 and 10) and two choices of σ (0.1 and 0.3).

Simulation Designs

- To perform an unbiased comparison of the three tree removal measures, we simulated three independent data sets
 - The training set is used to train the initial random forest
 - The execution set is used to delete trees from the initial forest to produce sub-forests
 - The evaluation set is used to evaluate the prediction performance of the sub-forests
- The generation and use of these three data sets constituted one run of simulation, and we replicated 100 times.

Simulation Results

- Randomly selected one run of simulation and presented the stepwise change in the prediction performance in Figure 1.
- The “by prediction” method is preferable
 - It can identify a critical point during the tree removal process in which the performance of the sub-forest deteriorates very rapidly.
- The performance of the sub-forests may begin to improve before the critical point.

$$P\{M_i|y_i\} = \frac{1}{P\{y_i\}} \prod_j [P\{y_{ij}|c_{ij} = 0\}] P\{M_i\}$$

$$= \frac{P\{M_i\}}{P\{y_i\}} \prod_j [\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\}]$$

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$$P\{y_i\} = \prod_j [\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\}] P\{M_i\}$$

$$\frac{\partial}{\partial \beta} \log(P\{M_i|y_i\}) = -\frac{\partial}{\partial \beta} \log(P\{y_i\})$$

$$= [1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)]$$

$$= \sum_j \frac{1 - \gamma(y_{ij}) - \gamma(y_{ij}, 0)}{P\{M_{ij}\}}$$

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$$P\{M_i|y_i\} = \frac{1}{P\{y_i\}} \prod_j [P\{y_{ij}|c_{ij} = 0\}] P\{M_i\}$$

$$= \frac{P\{M_i\}}{P\{y_i\}} \prod_j [\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\}]$$

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$$= [1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)]$$

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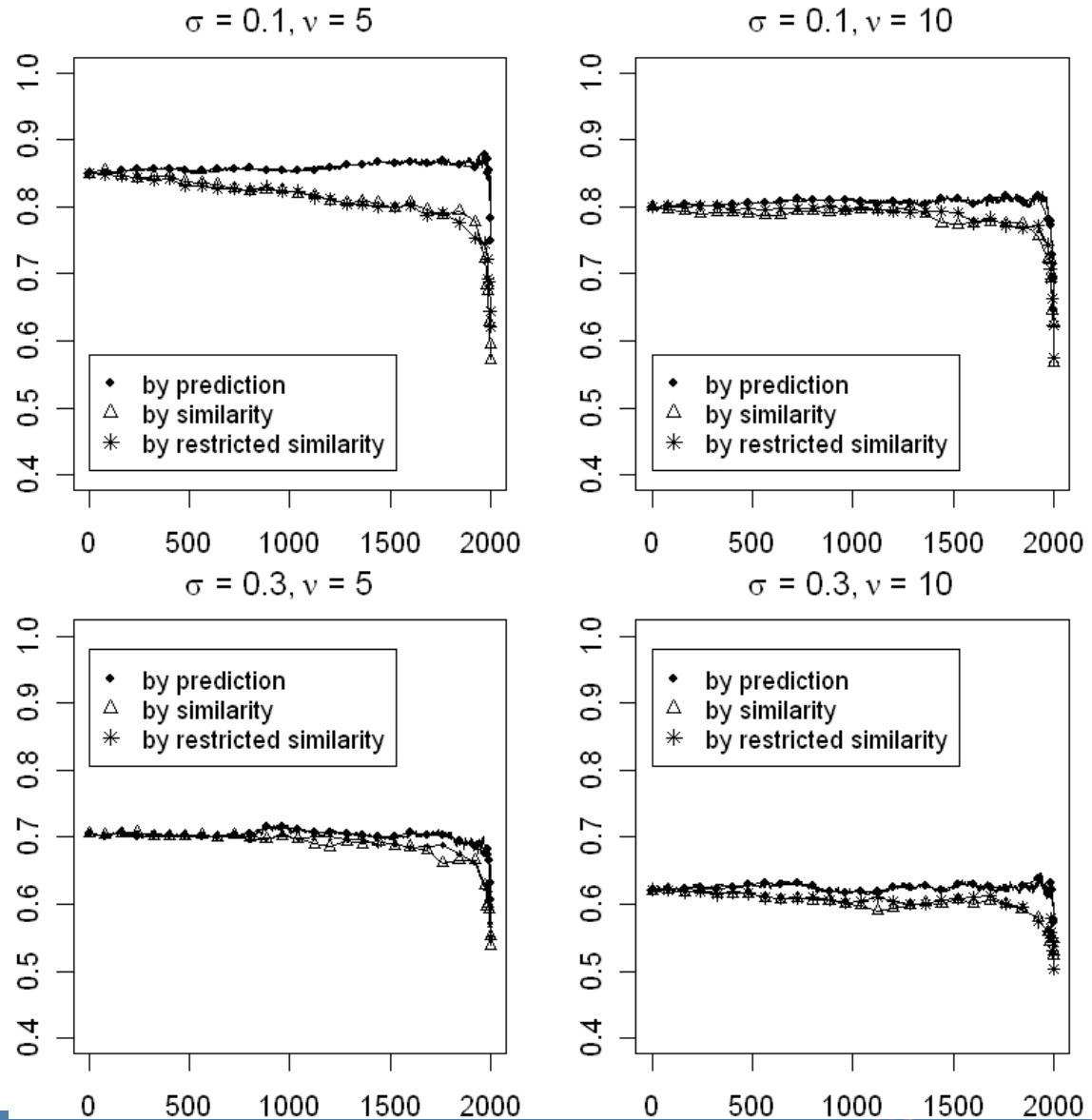
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Prediction performance of sub-forests produced from different datasets and methods



Simulation Results

- Figure 2 displays a summary plot of prediction performance using the results in five randomly selected runs.
- Although the variation of the trajectories is notable, the sizes of the optimal subforests are within a reasonable range (11-36) for the “by prediction” method.

$$P\{M_i|y_i\} = \frac{1}{P\{y_i\}} \prod_j [P\{y_{ij}|c_{ij} = 0\} P\{\dots\}]$$

$$= \frac{P\{M_i\}}{P\{y_i\}} \prod_j [\pi(\beta; y_{ij}, 0) P\{\dots\}]$$

$$P\{y_{ij} = k | c_{ij} = c\} = \gamma(\beta; k, c)$$

$$K - 1, \gamma(\beta, 0, c) = 0, \text{ and } \gamma(\beta, K, c) = 1$$

$$P\{y_i\} = \prod_j [P\{y_{ij}|c_{ij} = 0\} P\{\dots\}]$$

$$= \prod_j [\pi(\beta; y_{ij}, 0) P\{\dots\}]$$

able to see that $(\partial/\partial\beta)\pi(\beta; k, c) = c$

$$\log(P\{M_i|y_i\}) = \frac{\partial}{\partial\beta} \log(P\{y_i\})$$

the number of runs β

$$\frac{\partial}{\partial\beta} \log(P\{M_i|y_i\}) = \sum_j \frac{\partial}{\partial\beta} \log(\pi(\beta; y_{ij}, 0) P\{\dots\})$$

$$= [1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)]$$

venience, we drop the two irrelevant

$$g(P\{M_i|y_i\})|_{\beta=0} = \sum_j [1 - \gamma(y_{ij}) - \gamma(y_{ij}, 0)]$$

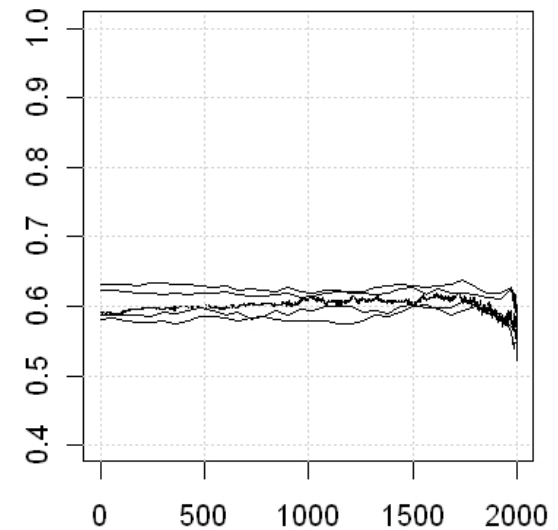
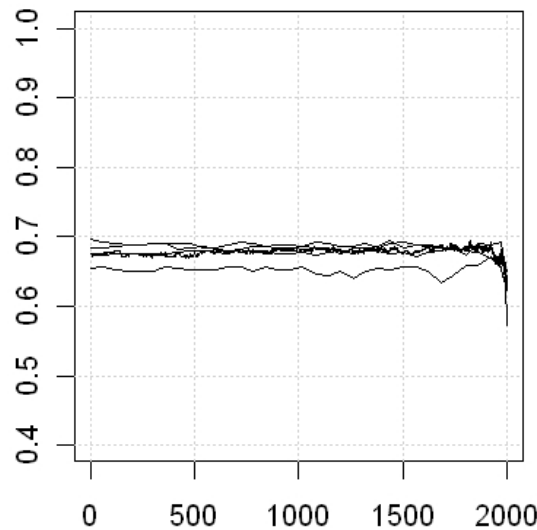
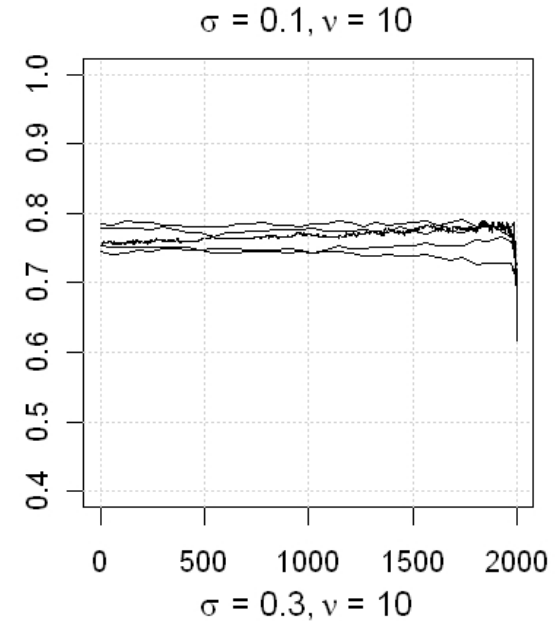
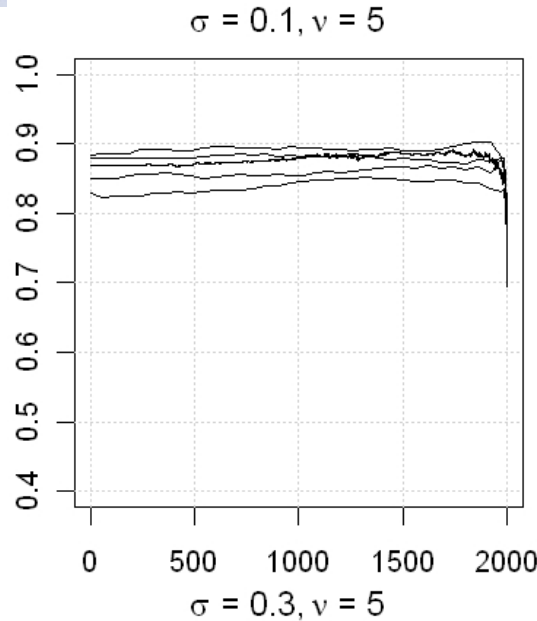
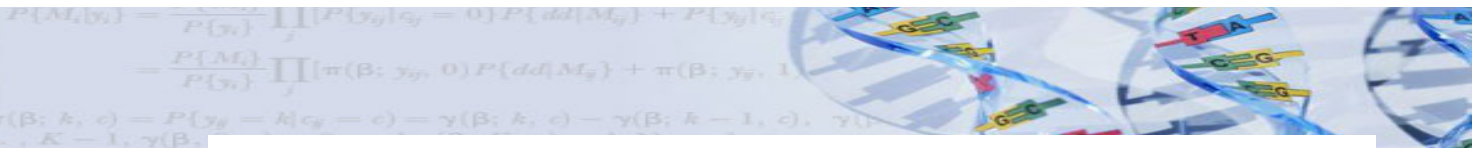
$$= \sum_j \frac{1 - \gamma(y_{ij}) - \gamma(y_{ij}, 0)}{P\{M_{ij}\}}$$

the coefficient of linkage disequilibrium

$$P\{AA\} - P\{dd, AA\} - P\{AA\}P\{DE\}$$

, AA} - P{dd, AA} - P{AA}P{DE

Performance trajectory of the “by prediction” method using the results in five randomly selected runs for four data sets.



$$P\{M_i|y_i\} = \frac{1}{P\{y_i\}} \prod_j [P\{y_{ij}|c_{ij} = 0\}] P\{dd[M_{ij}] + P\{y_{ij}|c_{ij} = 1\}$$

$$= \frac{P\{M_i\}}{P\{y_i\}} \prod_j [\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\} + \pi(\beta; y_{ij}, 1) P\{c_{ij} = 1\}]$$

$$P\{y_{ij} = k|c_{ij} = c\} = \gamma(\beta; k, c) - \gamma(\beta; k-1, c), \quad \gamma(\beta; 0, c) = 0, \text{ and } \gamma(\beta; K, c) = 1.$$

$$P\{y_i\} = \prod_j [P\{y_{ij}|c_{ij} = 0\}] P\{c_{ij} = 0\} + \prod_j [\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\} + \pi(\beta; y_{ij}, 1) P\{c_{ij} = 1\}]$$

to see that $(\partial/\partial\beta)\pi(\beta; k, c) = c - \gamma(\beta; k, c) + \gamma(\beta; k-1, c)$.

$$\log(P\{M_i|y_i\}) = -\frac{\partial}{\partial\beta} \log(P\{y_i\}) + \sum_j \frac{\partial}{\partial\beta} \log[\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\} + \pi(\beta; y_{ij}, 1) P\{c_{ij} = 1\}]$$

Under the null hypothesis that $\beta = 0$, we have $\gamma(0; k, c) = c$.

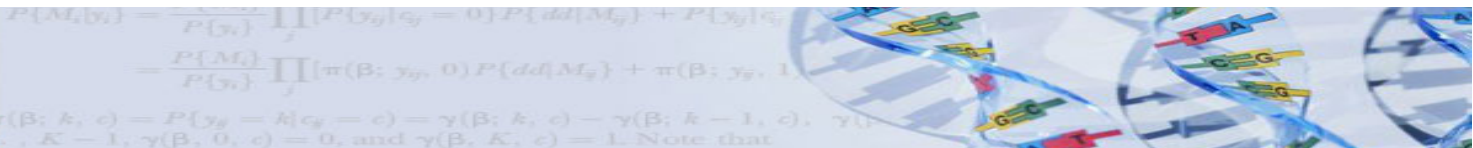
$$\frac{\partial}{\partial\beta} \log[\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\}] = [1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)]$$

$$\frac{\partial}{\partial\beta} \log P\{y_i\}|_{\beta=0} = \sum_j [1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)]$$

For convenience, we drop the two irrelevant terms in the sum.

$$\log(P\{M_i|y_i\})|_{\beta=0} = \sum_j [1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)] - \sum_j \frac{1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)}{P\{M_{ij}\}}$$

The coefficient of linkage disequilibrium is $D = P\{AA\} - P\{dd, AA\} - P\{AA\}P\{DE\}$.



The medians of the numbers of trees in the optimal sub-forests in 100 replications.

σ	ν	
	5	10
0.1	20(13, 29)	31(20, 40)
0.3	22(15, 32)	18(11, 37)

Simulation Designs

- In practice, we generally have one data set only.
- May not have the execution and evaluation data sets as in previous simulation.
- How do we select the optimal sub-forest with only one data set?

Simulation Designs

- Considered four bootstrap-based approaches and examined them in simulated data sets.
- We have the “golden” standard to be compared with in the simulated data set.

$$P\{M_i|y_i\} = \frac{P\{M_i\}}{P\{y_i\}} \prod_j [P\{y_{ij}|c_{ij} = 0\}]$$

$$= \frac{P\{M_i\}}{P\{y_i\}} \prod_j [\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\}]$$

$$P\{y_{ij} = k|c_{ij} = c\} = \gamma(\beta; k, c)$$

$$K - 1, \gamma(\beta, 0, c) = 0, \text{ and } \gamma(\beta, K, c) = 1.$$

$$P\{y_i\} = \prod_j [P\{y_{ij}|c_{ij} = 0\}]$$

$$= \prod_j [\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\}]$$

able to see that $(\partial/\partial\beta)\pi(\beta; k, c) = c - \gamma(\beta; k, c)$

$$\log(P\{M_i|y_i\}) = -\frac{\partial}{\partial\beta} \log(P\{y_i\})$$

$$+ \sum_j \frac{\partial}{\partial\beta} \log[\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\}]$$

the null hypothesis that $\beta = 0$, we have

$$\frac{\partial}{\partial\beta} \log[\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\}]$$

$$= [1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)]$$

$$\frac{\partial}{\partial\beta} \log P\{y_i\}|_{\beta=0} = \sum_j [1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)]$$

convenience, we drop the two irrelevant terms

$$\log(P\{M_i|y_i\})|_{\beta=0} = \sum_j [1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)]$$

$$= \sum_j \frac{1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)}{P\{M_{ij}\}}$$

the coefficient of linkage disequilibrium

$$D = P\{AA\} - P\{dd, AA\} - P\{AA\}P\{DE\}$$

Simulation Designs

- After constructing an initial forest using the whole data set as the training data set
 - use one bootstrap data set for execution and the out-of-bag (oob) samples for evaluation.
 - use the oob samples for both execution and evaluation.
 - use the bootstrap samples for both execution and evaluation.
 - re-draw bootstrap samples for execution and re-draw bootstrap samples for evaluation.

Simulation Results

- Figure 3 compares the performance of the four bootstrap-based approaches in the four simulation data sets.
- The comparison is based on the average performance in 100 runs.

$$P\{M_i|y_i\} = \frac{1}{P\{y_i\}} \prod_j [P\{y_{ij}|c_{ij} = 0\} P\{c_{ij} = 0\}]$$

$$= \frac{P\{M_i\}}{P\{y_i\}} \prod_j [\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\}]$$

$$P\{y_i\} = \prod_j [P\{y_{ij}|c_{ij} = 0\} P\{c_{ij} = 0\} + P\{y_{ij}|c_{ij} = 1\} P\{c_{ij} = 1\}]$$

$$P\{M_i\} = \prod_j [\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\}]$$

A performance summary plot of the “by prediction” method

$$\frac{\partial}{\partial \beta} \log P\{y_i\} = \sum_j [1 - \gamma(y_{ij}) - \gamma(y_{ij}, 1) - \gamma(y_{ij}, 0)]$$

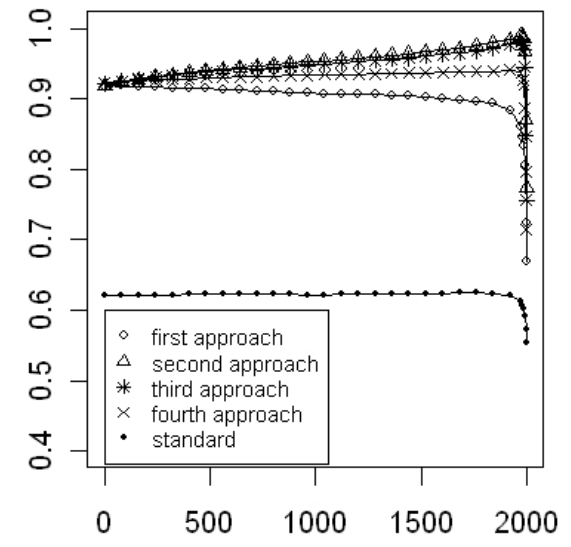
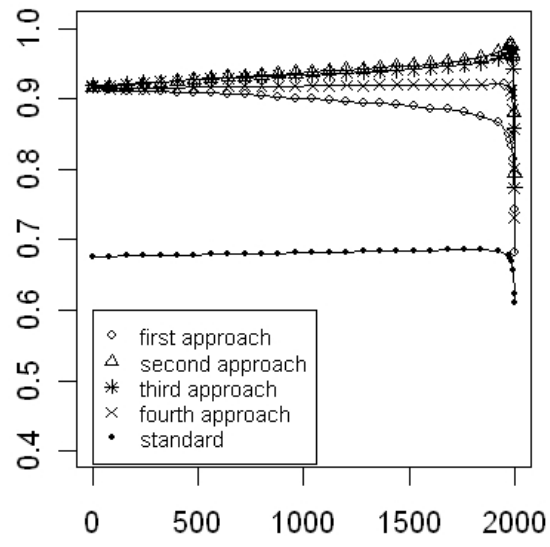
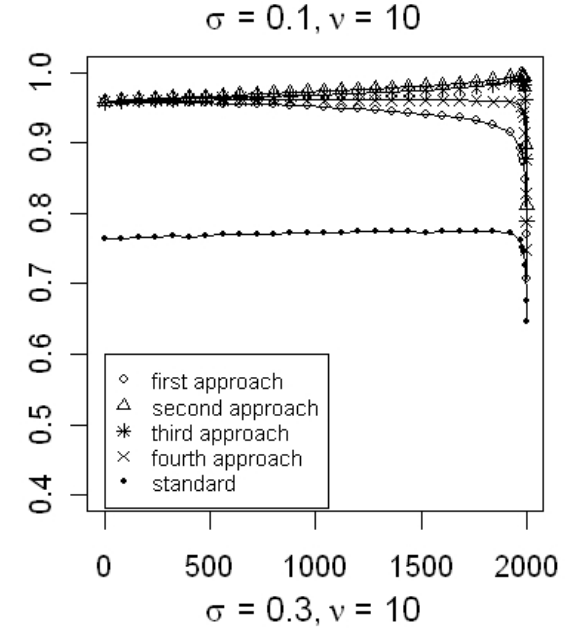
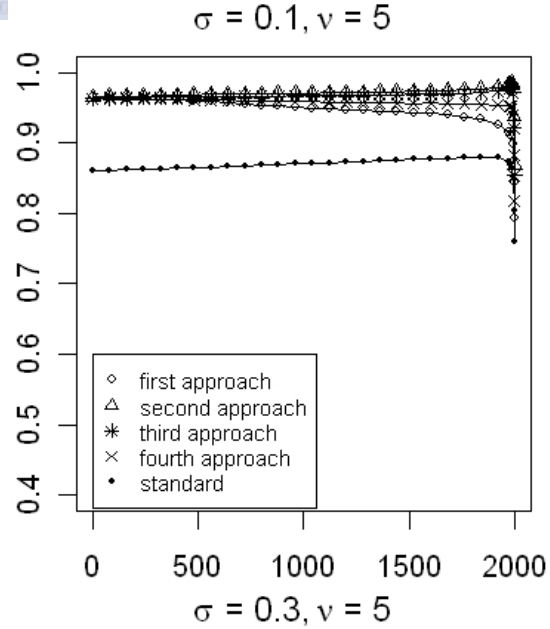
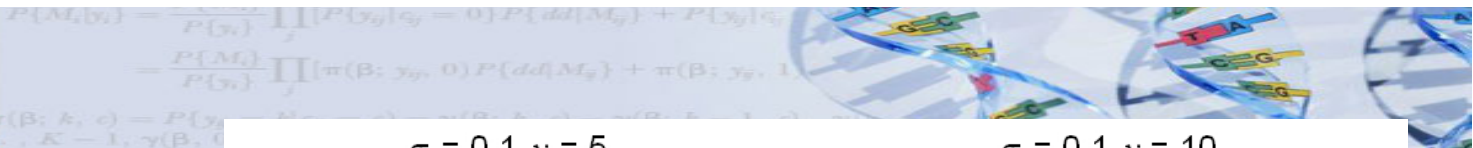
convenience, we drop the two irrelevant

$$\frac{\partial}{\partial \beta} \log P\{M_i|y_i\} = \sum_j [1 - \gamma(y_{ij}) - \gamma(y_{ij}, 1) - \gamma(y_{ij}, 0)]$$

$$= \sum_j \frac{1 - \gamma(y_{ij}) - \gamma(y_{ij}, 1) - \gamma(y_{ij}, 0)}{P\{M_i|y_i\}}$$

the coefficient of linkage disequilibrium

$$P\{AA\} - P\{dd, AA\} - P\{AA\}P\{DE\}$$



Simulation Results

- The performance trajectories of the four bootstrap-based approaches may not overlap with the “golden” standard.
- For the selection of the optimal subforest, the similarity among the trajectories is most relevant, because it could lead to the same or similar subforest.

Simulation Results

- In Figure 4, we examined the correlation between the original (the “golden” standard) trajectory and each of the four bootstrap approaches.

$$P\{M_i|y_i\} = \frac{1}{P\{y_i\}} \prod_j [P\{y_{ij}|c_{ij} = 0\}] P\{c_{ij} = 0\}$$

$$= \frac{P\{M_i\}}{P\{y_i\}} \prod_j [\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\}]$$

$$P\{y_i = k, c_i = c\} = P\{y_{ij} = k | c_{ij} = c\} = \gamma(\beta; k, c)$$

$$K - 1, \gamma(\beta, 0, c) = 0, \text{ and } \gamma(\beta, K, c) = 1. \text{ Note that}$$

$$P\{y_i\} = \prod_j [P\{y_{ij}|c_{ij} = 0\}] P\{c_{ij} = 0\}$$

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able to see that $(\partial/\partial\beta)\pi(\beta; k, c) = c - k$

$$\log(P\{M_i|y_i\}) = -\frac{\partial}{\partial\beta} \log(P\{y_i\})$$

$$+ \sum_j \frac{\partial}{\partial\beta} \log[\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\}]$$

the null hypothesis that $\beta = 0$, we have

$$\frac{\partial}{\partial\beta} \log[\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\}]$$

$$= [1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)]$$

$$\frac{\partial}{\partial\beta} \log P\{y_i\}|_{\beta=0} = \sum_j [1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)]$$

convenience, we drop the two irrelevant terms

$$\log(P\{M_i|y_i\})|_{\beta=0} = \sum_j [1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)]$$

$$= \sum_j \frac{1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)}{P\{M_{ij}\}}$$

the coefficient of linkage disequilibrium

$$D = P\{AA\} - P\{dd, AA\} - P\{AA\}P\{DE\}$$

$$P\{M_i|y_i\} = \frac{1}{P\{y_i\}} \prod_j [P\{y_{ij}|c_{ij} = 0\} P\{c_{ij} = 0\}]$$

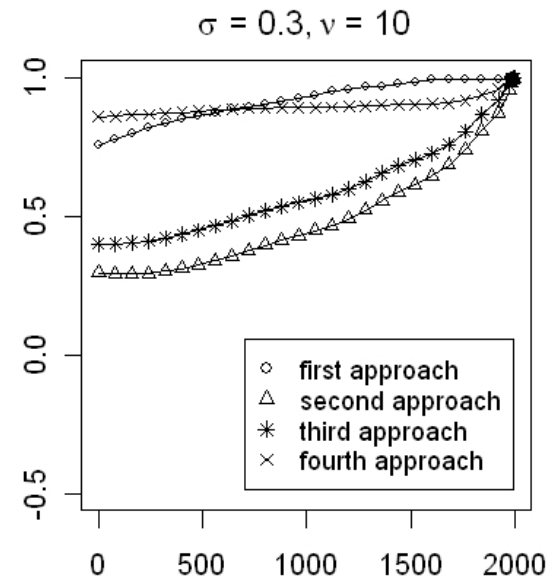
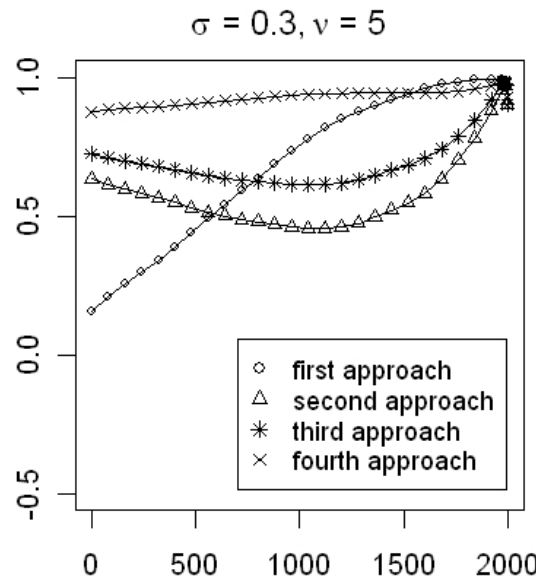
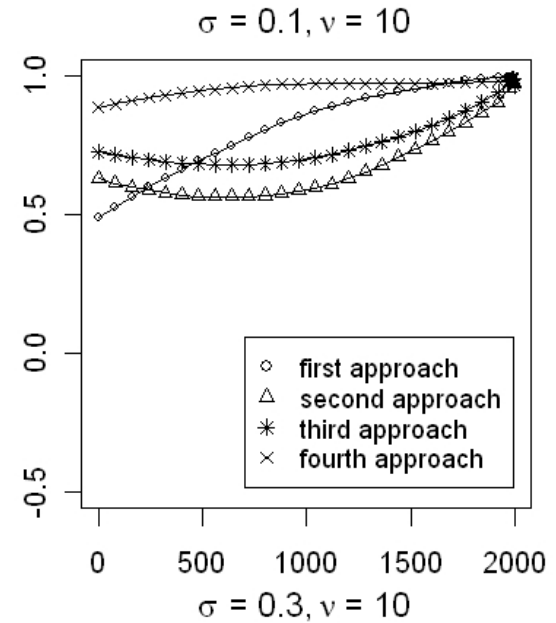
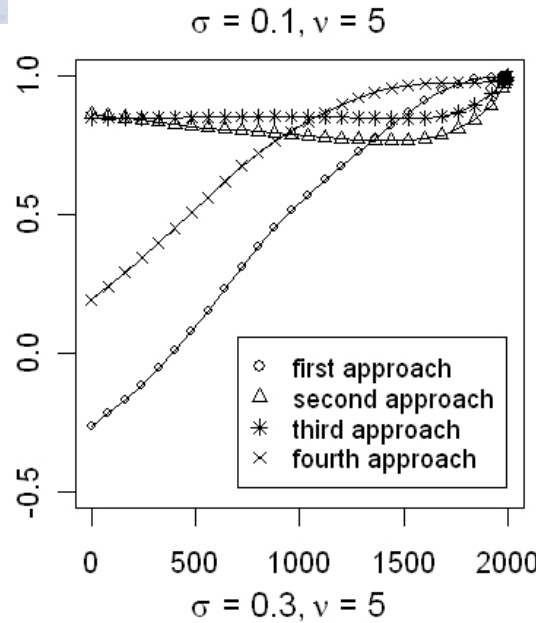
$$= \frac{P\{M_i\}}{P\{y_i\}} \prod_j [\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\}]$$

$$P\{y_i = k, c_i = \gamma\} = P\{y_i = k | c_i = \gamma\} P\{c_i = \gamma\}$$

$$= \frac{P\{M_i\}}{P\{y_i\}} \prod_j [\pi(\beta; y_{ij}, 0) P\{c_{ij} = 0\}] P\{c_i = \gamma\}$$

$$P\{y_i\} = \prod_j [P\{y_{ij}|c_{ij} = 0\} P\{c_{ij} = 0\}]$$

The correlation between the performance trend by each of the four bootstrap strategies and the "standard" curve



Simulation Results

- Using the bootstrap samples for execution and the oob samples for evaluation is an effective sample-reuse approach to selecting the optimal subforest.

Application

- Dataset

- the microarray data set of a cohort of 295 young patients with breast cancer, containing expression profiles from 70 previously selected genes.

- previously studied by van de Vijver *et al.*

- The responses of all patients are defined by whether the patients remained disease-free five years after their initial diagnoses or not.

Application

- Method used

- The “by prediction” measure
- The original data set to construct an initial forest
- A bootstrap data set for execution
- The oob samples for evaluation.

- The procedure is replicated for a total of 100 times.

- The oob error rate is used to compare the performance of the initial random forest and the optimal sub-forest.
- The sizes of the optimal sub-forests fall in a relatively narrow range, of which the 1st quartile, the median, and the 3rd quartile are 13, 26 and 61, respectively.

Application

- The smallest optimal sub-forest in the 100 repetitions with the size of 7 is selected.
- As a benchmark, we used the 70-gene classifier proposed by Vijver, *et al.*

Application

• Table 2 presents the misclassification rates based on the oob samples.

– The initial forest and the optimal sub-forest achieve almost the same level of performance accuracy.

– The 70-gene classifier has an out-of-bag error rate which is much higher than those of the forests.

Comparison of prediction performance of the initial random forest, the optimal sub-forest, and a previously established 70-gene classifier

Method	Error rate	True		
		Predicted	Good	Poor
Random Forest	26.0%	Good	141	17
		Poor	53	58
Sub-forest	26.0%	Good	146	22
		Poor	48	53
70-gene Classifier	35.3%	Good	103	4
		Poor	91	71

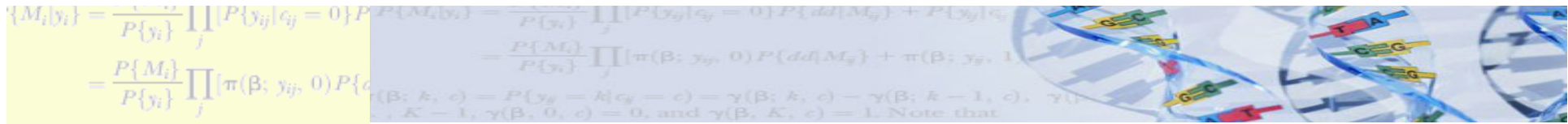
Application

- Main motivation

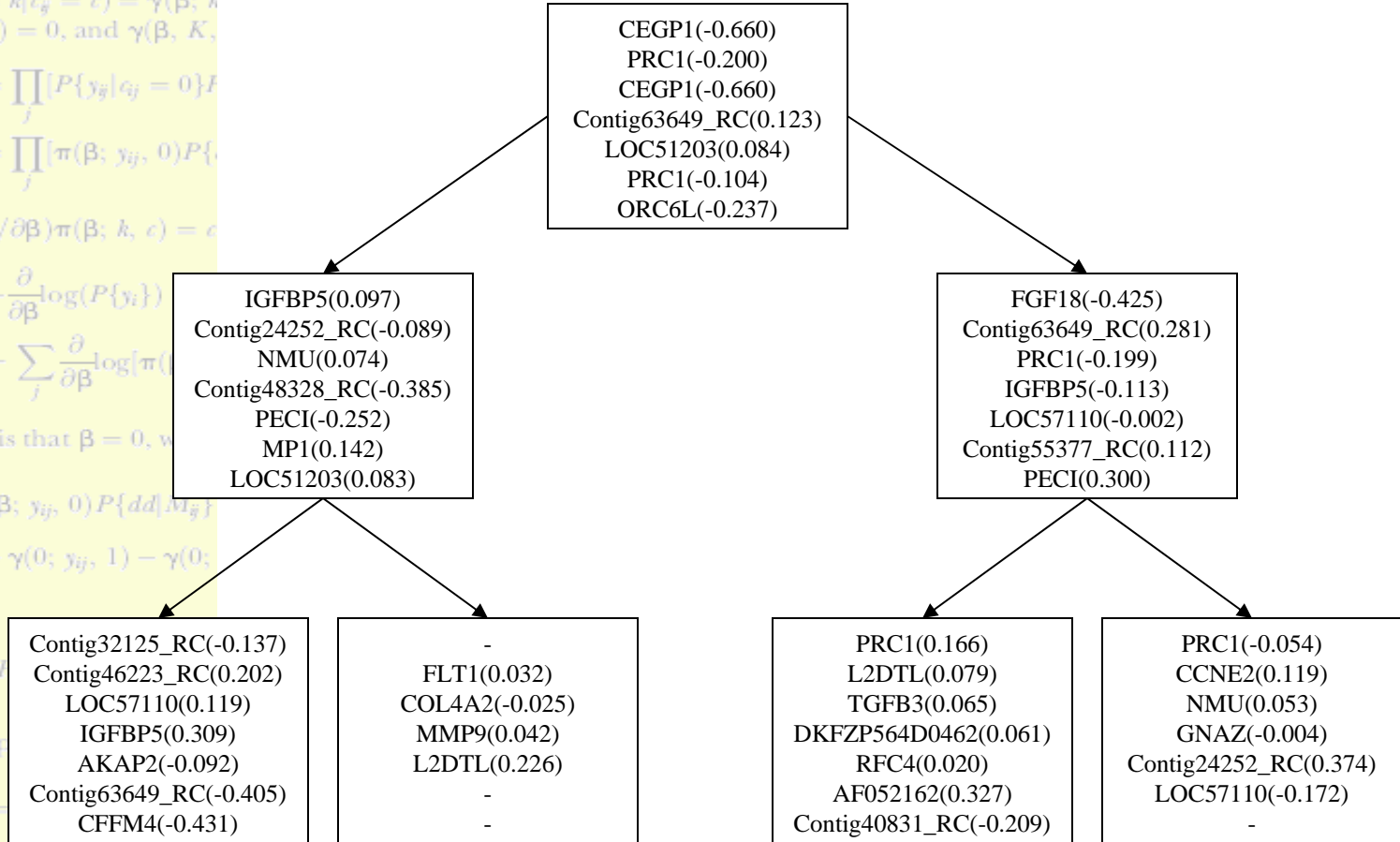
- seek the smallest possible forest to enable us to examine the forest.

- Figure 5 displays the most critical part (the top three layers) of the optimal sub-forest consisting of the seven trees.

- The selected genes are quite diverse and unique.



$P\{M_i|y_i\} = \frac{1}{P\{y_i\}} \prod_j [P\{y_{ij}|c_{ij} = 0\}] P\{dd[M_i]\}$
 $= \frac{P\{M_i\}}{P\{y_i\}} \prod_j [\pi(\beta; y_{ij}, 0) P\{dd[M_i]\}]$
 $P\{y_i\} = \prod_j [P\{y_{ij}|c_{ij} = 0\}] P\{dd[M_i]\}$
 $= \prod_j [\pi(\beta; y_{ij}, 0) P\{dd[M_i]\}]$
 $\log(P\{M_i|y_i\}) = -\frac{\partial}{\partial \beta} \log(P\{y_i\})$
 $+ \sum_j \frac{\partial}{\partial \beta} \log[\pi(\beta; y_{ij}, 0) P\{dd[M_i]\}]$
 $\frac{\partial}{\partial \beta} \log[\pi(\beta; y_{ij}, 0) P\{dd[M_i]\}] = [1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)]$
 $\frac{\partial}{\partial \beta} \log(P\{M_i|y_i\})|_{\beta=0} = \sum_j \frac{1 - \gamma(y_{ij}) - \gamma(y_{ij}, 0)}{1 - \gamma(y_{ij}, 0)}$



The top three layers of the optimal sub-forest consisting of seven trees

Conclusion

- It is possible to construct a highly accurate random forest consisting of a manageable number of trees.

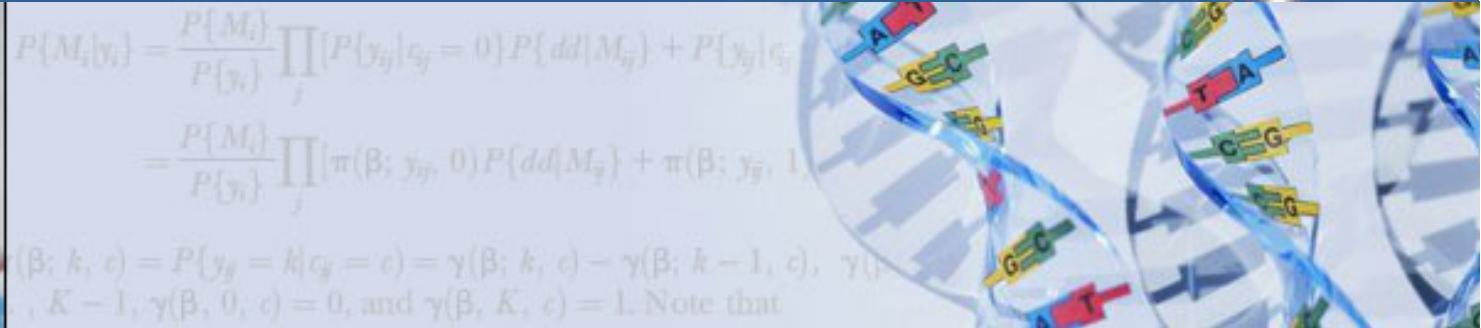
- the size of the optimal sub-forest is in the range of tens
- some sub-forests can even over-perform the original forest in terms of prediction accuracy

- The key advantage

- the ability to examine and present the forests.

- The limitation

- future samples and studies are needed to evaluate the performance of the forest-based classifiers.



$$P\{M_i|y_i\} = \frac{P\{M_i\}}{P\{y_i\}} \prod_j [P\{y_{ij}|c_j = 0\}P\{dd|M_j\} + P\{y_{ij}|c_j = 1\}P\{dd|M_j\}]$$

$$= \frac{P\{M_i\}}{P\{y_i\}} \prod_j [\pi(\beta; y_{ij}, 0)P\{dd|M_j\} + \pi(\beta; y_{ij}, 1)P\{dd|M_j\}]$$

$\pi(\beta; k, c) = P\{y_j = k|c_j = c\} = \gamma(\beta; k, c) - \gamma(\beta; k-1, c)$, $\gamma(\beta; 0, c) = 0$, $\gamma(\beta; K-1, c) = 1$, $\gamma(\beta, 0, c) = 0$, and $\gamma(\beta, K, c) = 1$. Note that

able to see that $(\partial/\partial\beta)\pi(\beta; k, c) = c$

$$\log(P\{M_i|y_i\}) = -\frac{\partial}{\partial\beta}\log(P\{y_i\})$$

$$+ \sum_j \frac{\partial}{\partial\beta}\log[\pi(\beta; y_{ij}, 0)P\{dd|M_j\} + \pi(\beta; y_{ij}, 1)P\{dd|M_j\}]$$

Under the null hypothesis that $\beta = 0$, we have

$$\frac{\partial}{\partial\beta}\log[\pi(\beta; y_{ij}, 0)P\{dd|M_j\}] = [1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)]$$

$$\frac{\partial}{\partial\beta}\log P\{y_i\}|_{\beta=0} = \sum_j [1 - \gamma(0; y_{ij}, 1) - \gamma(0; y_{ij}, 0)]$$

For convenience, we drop the two irrelevant terms

$$\log(P\{M_i|y_i\})|_{\beta=0} = \sum_j [1 - \gamma(y_{ij}, 1) - \gamma(y_{ij}, 0)]$$

$$= \sum_j \frac{1 - \gamma(y_{ij}, 1) - \gamma(y_{ij}, 0)}{P\{M_j\}}$$

is the coefficient of linkage disequilibrium

$$D = P\{AA\} - P\{dd, AA\} - P\{AA\}[P\{DE\} - P\{DE\}]$$

Thank You!